

11. Nandagopalan, S., Adiga, B. S. and Deepak, N., A universal model for content-based image retrieval. *Int. J. Electr. Comput. Eng.*, 2009, **4**.
12. Malik, F. and Baharudin, B. B., Feature analysis of quantized histogram color features for content-based image retrieval based on Laplacian filter. In International Conference on System Engineering and Modeling. IACSIT Press, Singapore, 2012, vol. 34.
13. Patil, M. P. and Kolhe, S. R., Automatic image categorization and annotation using K-NN for corel dataset. *Adv. Comput. Res.*, **4**, 108–112.
14. Chapelle, O., Haffner, P. and Vapnik, V. N., Support vector machines for histogram-based image classification. *IEEE Trans. Neural Networks*, 1999, **10**, 1055–1064.
15. Chen, Y. and Wang, J. Z., A region-based fuzzy feature matching approach to content-based image retrieval. *IEEE Trans. Pattern Anal. Mach. Intell.*, 2002, **24**, 1252–1267.
16. Chang, C.-C. and Lin, C.-J., LIBSVM: a library for support vector machines, 2001; Software available at <http://www.csie.ntu.edu.tw/~cjlin/libsvm>
17. Li, J., Wang, J. Z. and Wiederhold, G., IRM: integrated region matching for image retrieval. Proceedings of the 8th ACM International Conference on Multimedia, Los Angeles, CA, USA, 2000, pp. 147–156.

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## Application of fractal geometry in determining optimal quadrat size for vegetation sampling

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**Geometry in ecological patterns of landscape and vegetation is not truly fractal, and varies across a range of scales, whereas fractal geometry provides tools for predicting and describing ecological patterns. In this study, fractal analysis is used to assess presence of pseudo random quadrats or spatial dependence which hamper generality and performance of classical inferential statistics. Fractal dimension (FD) as a function of scale is used to determine quadrat size which eliminates spatial dependence. The semivariograms are plotted with fractograms to correlate structures of spatial dependence of the properties studied. The use of FD as a degree of spatial dependence of variables is the basis of applications of fractals.**

**Keywords:** Ecological patterns, fractal geometry, quadrat size, spatial dependence, vegetation sampling.

A significant challenge encountered in plant ecological studies is vegetation sampling<sup>1,2</sup>. Researchers worldwide

have analysed ecological attributes (species diversity, richness, dominance, etc.) of vegetation using random or stratified random sampling or by laying transects across some gradient<sup>3–6</sup>. Ecologists have used larger contiguous area and researchers have also designed certain plots as ‘long-term ecological plots’ or ‘permanent dynamic plots’ to monitor variability in species characteristics in spatio-temporal domain<sup>7–15</sup>. Whichever method is adopted, sampling is always a time-consuming process. Also, the size of the plot or quadrat that is used as the basic unit of sampling, varies depending on the type of vegetation and area covered. Though significant variation exists in the ecological patterns captured by random method (usually high and diverse) compared with large-area contiguous plots, these are basically used to understand the behavioural patterns of the species in contiguous scale<sup>11,16–19</sup>.

Studying the large-area plots (which range between 1 and 50 ha), researchers have subdivided the entire plot into smaller units for better and quick sampling. The size of the smaller units are 1 m × 1 m, 10 m × 10 m, 30 m × 30 m or sometimes circular plots with varied dimensions<sup>2,19–22</sup>. Within sub-units, ecologists study characteristic features of a species and its population or general diversity patterns, and compare the changing attributes across the quadrats conceptualizing the pattern at higher scale<sup>23</sup>. But when comparisons are made between neighbouring or adjacent quadrats, probability of variation is low as it lies in the same homogenous conditions – may be precipitation, edaphic, sometimes topography. This indicates greater similarity in two closely spaced quadrats compared to those that are separated by larger distances. These samples may be referred to as pseudo replicates, violating the most important assumption of classical inferential statistics that the samples are spatially independent<sup>24</sup>.

A homogeneous distribution is one that remains similar on repeated sub-division<sup>25</sup>. The arrangement or ordering of data as a function of location is called spatial autocorrelation of the function and the range of spatial scales in which spatial autocorrelation exists is called spatial dependence<sup>26</sup>. Avoiding spatially dependent quadrats (pseudo replicates), that do not contribute significant changes in any ecological property is necessary, to improve the performance of classical inferential statistics, as the existence of spatial dependence hampers the generality of results and overall performance of classical inferential statistics<sup>27</sup>.

The concepts of fractal geometry can suggest a better statistically rectified sampling scheme, which eliminates the problem of spatial dependence among the quadrats<sup>25</sup>. Optimal quadrat sizes for homogeneous or spatially independent distribution can be determined by using the methods of fractal analysis on the data. The quadrats of suggested sizes will be spatially independent and thus independent of the distances by which they are separated.

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Fractals<sup>28-41</sup> essentially have a non-integer dimension, unlike the corresponding Euclidean form, and can be described by fractal dimension (FD), which can be used to compare and categorize fractals<sup>40,42-44</sup>. A non-integer dimension implies that the fractal has a dimension different from the space in which it resides<sup>45</sup>. The FD has also been widely used as a measure of the space-filling ability of a pattern<sup>42,46</sup>.

A perfectly random distribution will be homogeneous and have a FD of two<sup>25,45</sup>. So by dividing the sample space into quadrats of size at which the FD is two, we can have a distribution which will be homogenous.

In the perspective of these concepts, an attempt was made to suggest appropriate quadrat sizes for homogeneous distribution best suited for the methods of classical statistics, using FD analysis and to suggest appropriate spacing between the quadrats, restricting the number of quadrats to be sampled to a lower value. The study further aims, as a secondary objective, at drawing inference regarding the variation and correlation among various properties studied in the contiguous plot.

We have used three contiguous plot data, each of size 3 ha sampled in three different forest types, viz. evergreen (EG), semi-evergreen (SEG) and moist deciduous (MD) forests of Northern Andaman Islands. Intact, isolated plots were selected and each 3 ha plot was divided into 30 quadrats of 0.1 ha size (32 m × 32 m). In each quadrat primary attributes of tree data such as their identification, number, height and girth at breast height (> 30 cm) were measured. The data were used to derive mean height, basal area, density, species richness (number of species in each quadrat) and diversity (Shannon index) in each plot. These five attributes are considered important in characterizing the vegetation patterns in contiguous plots.

Semivariogram is the basic unit of geostatistics which summarizes the variance observed in a dependent variable as a function of scale and is defined as

$$y(h) = \frac{2 \sum_{i=1}^{N(h)} (z(i) - z(i+h))^2}{2N(h)}, \quad (1)$$

where  $y$  is the semivariance at scale  $h$ ;  $z$  the dependent variable, at any point  $i$ , has the value  $z(i)$ ;  $z(i+h)$  the value of the dependent variable at a point separated from the point  $i$  by a distance  $h$  and  $N(h)$  is the number of points separated by distance  $h$ .

The properties with similar semivariograms have similar structure of spatial dependence<sup>26</sup>. The slope  $m$  of the double logarithmic plot of the semivariogram gives the FD of the distribution  $D$ . The slope of the double logarithmic semivariogram for every distance  $h$  is calculated as<sup>25</sup>

$$m = \frac{\log\{y(2h)\} - \log\{y(h)\}}{\log(2h) - \log(h)}. \quad (2)$$

This simplifies to,

$$m = \frac{\log\{y(2h)\} - \log\{y(h)\}}{\log(2)}. \quad (3)$$

This formula can be considered as the change in variance of a dependent variable as one doubles the scale. The FD is calculated by<sup>47,48</sup>

$$D = \frac{4-m}{2}, \quad (4)$$

and plotted as a function of scale that can be used to interpret the scale at which the distribution can be considered as homogeneous. The degree of spatial dependence across a range of sample scales is indicated by FD, whereas semivariograms indicate the structure of scale dependence<sup>27</sup>.

According to eq. (1), if  $z$  is a linear function of  $h$ , the semivariogram will be a parabola, as  $z(i) - z(i+h)$  is a linear function of  $h$ . A linear function squared is a parabola, the slope of the double logarithmic plot of a parabola is 2, which corresponds to a FD of 1. If the values of  $z$  in two near samples are no more or less different from two distant samples, in other words, if the distribution is perfectly homogeneous, the slope of the semivariogram will be zero, corresponding to a FD of 2. The double logarithmic semivariogram may not be linear, hence both  $m$  and  $D$  are not necessarily constant functions of scale<sup>25</sup>.

Further, fractal geometry, in particular the FD is not restricted between 1 and 2, i.e. just a function of position on a line, but can be readily extended to between 2 and 3, as a function of points on the plane<sup>25</sup>.

At the first instance, one cannot check for homogeneity on scales less than that of the length of a side of the quadrat, which is 30.62 m, because of the restriction imposed by the unavailability of data within a quadrat. Further the formula applied, being fairly accurate, calculates FD by taking the difference between a scale value and the double value. This restricts the upper bound of the scale in the fractogram to half its value used to plot the semivariogram. This in turn ensures that the size of a quadrat does increase more than three times the original, which is half the maximum scale in the semivariogram.

As is evident from Figure 1 a, the curve of semivariance versus scale is a parabola, the slope of double logarithmic plot of the parabola will be two, and the corresponding FD will be one, that is, strictly linear spatial dependence. FD is extremely close to one at all the scale points (Figure 1 b). The slight error is because input was generated only to five decimal points, as  $\sqrt[3]{2}$  is an irrational number; so more the decimal points, more accurate and close to one will be the result.

The difference in the values for a pair of closely spaced quadrats (lower scale) is more or less the same as compared to the difference in the values for a pair of quadrats

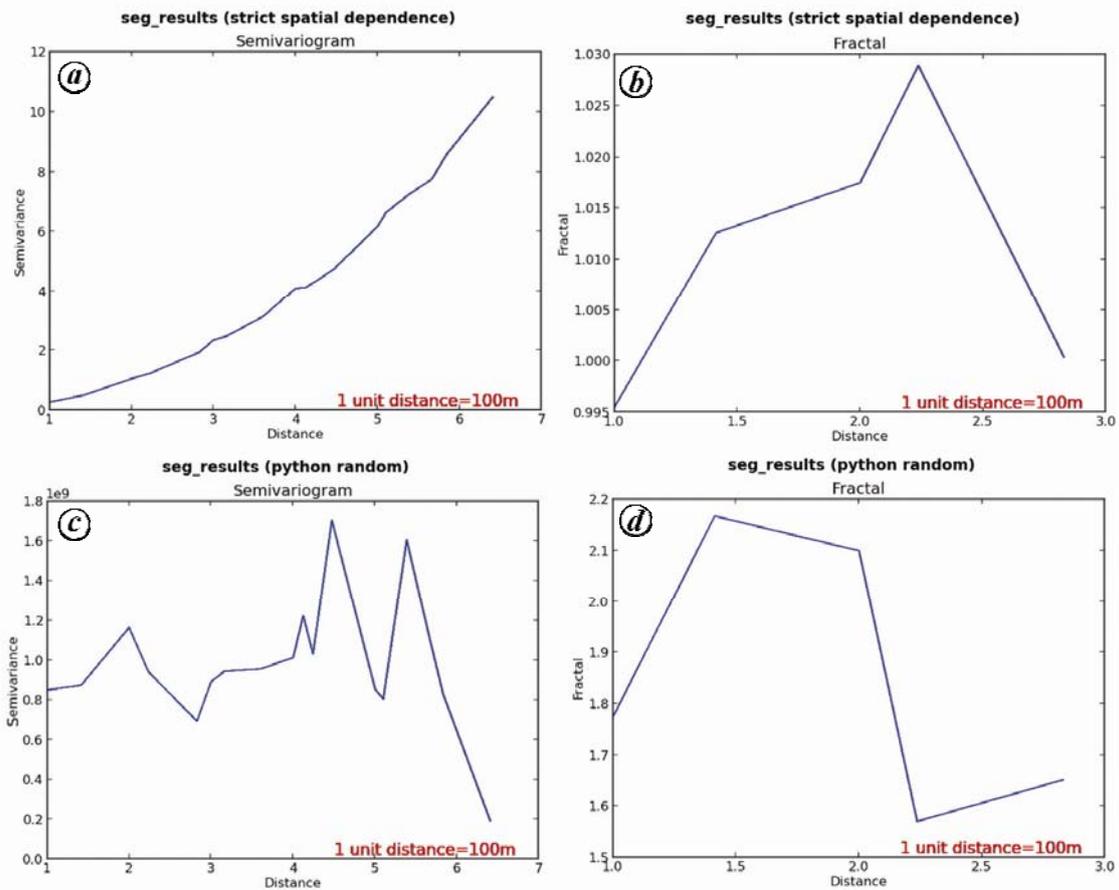


Figure 1. Verification of code.

placed at a larger distance (larger scale) (Figure 1 c). This implies that the slope of double logarithmic plot will be 0 and FD of the distribution will be 2 or close to 2 (random function of python is in effect pseudo random). In Figure 1 d, the FD is close to 2 at lower scale values and suffers a deviation due to pseudo random function of the python. A perfectly random distribution will result in a constant FD equal to 2, thus verifying the code and methodology applied.

Graphs like those of FD of height in EG and SEG forests, basal area in SEG forest (Figure 2) can be helpful in suggesting an ideal sampling scheme by choosing appropriate quadrat sizes, such that the quadrat delineates homogeneous vegetation, and quadrat spacing which allow spatial dependence<sup>26</sup>. The U-shaped section of the fractogram, with two limbs of U lying on  $d = 2$ , is necessary for suggesting such a scheme. We have suggested an ideal sampling scheme wherever we obtained such FD versus scale graphs. The quadrat side should be of size corresponding to where the left limb of U reaches  $d = 2$ . The quadrats should be separated by a minimum distance corresponding to where the right limb crosses  $d = 2$ . If the graph increases further and starts decreasing sharply, the corresponding distance should be the maximum distance between any two neighbouring quadrats.

The results are in effect invariant to white/coloured noise (of a constant amplitude) addition to the value of the function, shown below

$$z'(i) = z(i) + \text{noise}(0, 1) * \text{constant}. \quad (5)$$

The function  $\text{noise}(0, 1)$  is independent of  $i$  or the transect variable and follows a uniform random/Gaussian probability distribution depending upon the type of noise added. So,

$$z'(i) - z'(i + h) = z(i) - z(i + h) + \text{delta}, \quad (6)$$

where

$$\text{delta} = (\text{noise}^i(0, 1) - \text{noise}^{i+h}(0, 1)) * \text{constant}. \quad (7)$$

This study attempts to identify reliable plot size for field sampling, taken ideally for intact undisturbed forests and for such samples the value of the constant is small and hence the value of delta tends to zero. The value of the constant depends on sampling precision and is low for repeated and rigorous sampling, which is assumed. Addition of noise in general moves the FD towards that of a

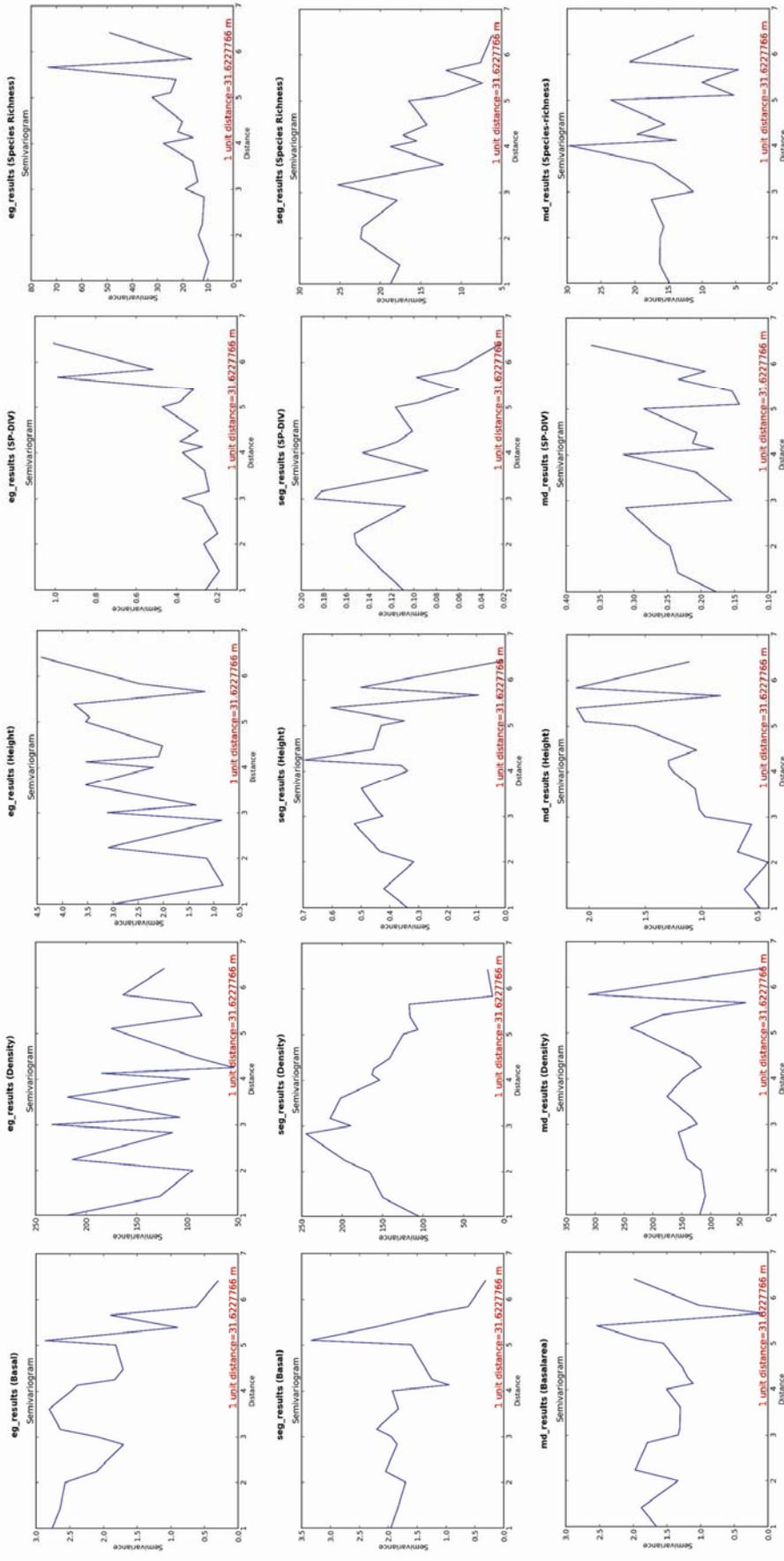


Figure 2. Semivariance versus scale/quadrant size/distance.

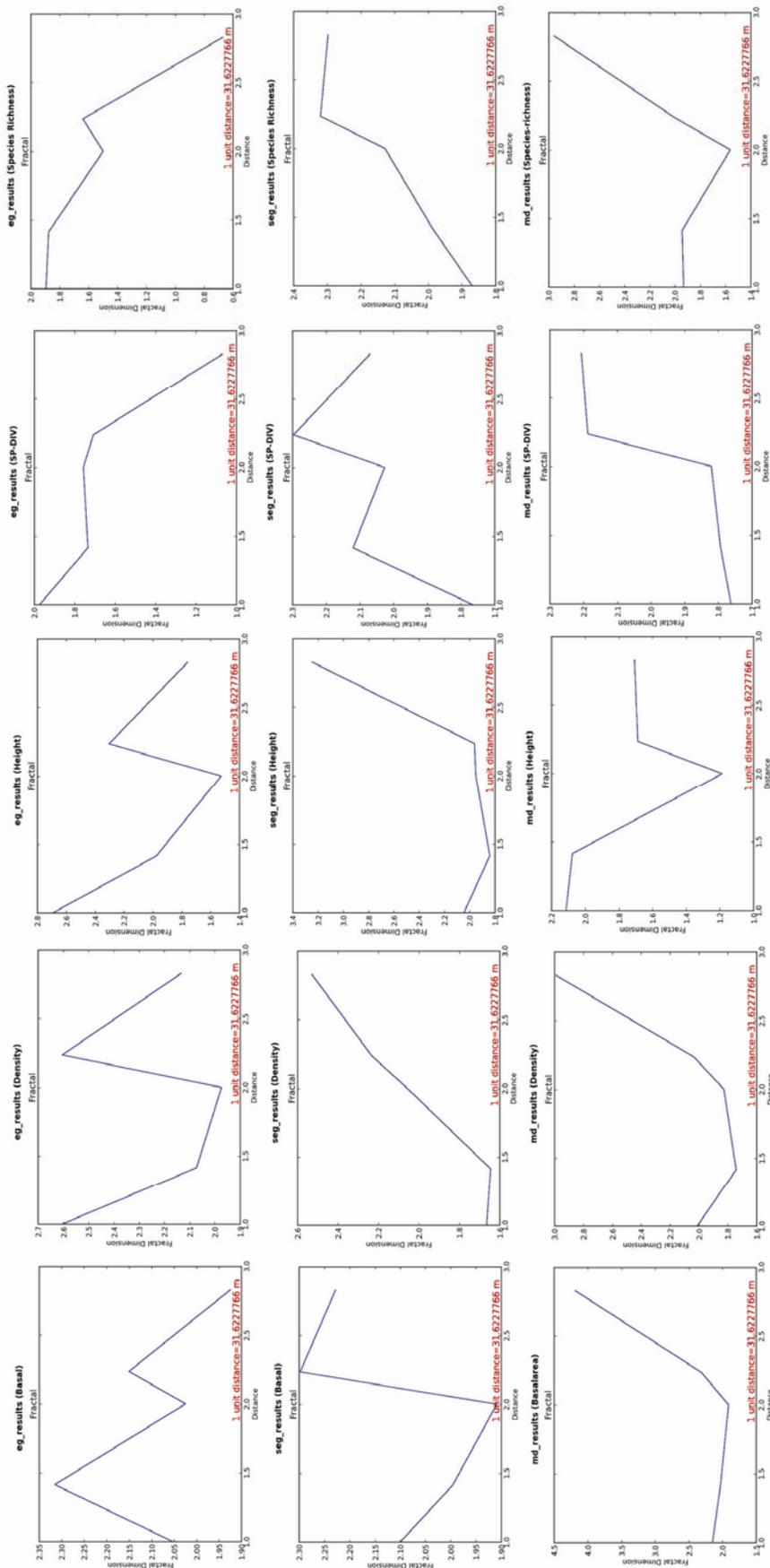


Figure 3. Fractal dimension versus scale/quadrant size/distance.

**Table 1.** Appropriate quadrat size

	Basal area	Density	Height	Species-dimension	Species diversity	Average for a forest type
Moist-deciduous	2.05	1.02	1.48	2.11	2.23	1.78
Semi-evergreen	2.05	1.90	1.08	1.27	1.45	1.55
Evergreen	2.00	1.83	1.42	1.00	1.00	1.45
Average for properties	2.03	1.58	1.33	1.46	1.56	

**Table 2.** Ideal sampling scheme

Plot	A	B	C
Evergreen → Height	1.42	2.13	2.27
Semi-evergreen → Basal area	1.4	2.05	2.25
Semi-evergreen → Height	1.08	2.26	–
Moist-deciduous → Basal area	1.55	2.05	–
Moist-deciduous → Density	1.02	2.19	–

homogenous distribution. We suggest quadrat size for homogenous distribution which will not be affected by added random noise. Further, all results are based on FD of the distribution, which is dependent on the slope of semivariance (see below), and for small values of the constant the slope is not affected. So in general the suggested quadrat size will not be affected by environmental disturbance.

These values can be used to study specific properties of specific forest types with sufficient generality<sup>25</sup>.

Figures 2 and 3 show the semivariance versus scale (quadrat size) and FD versus scale results respectively, for different properties for three forest types, viz. EG, SEG and MD.

In general, FD obtained was more close to 2 than 1, which suggests that vegetation has low spatial dependence. Nevertheless, the spatial dependence, even if it is low, hampers the performance of classical inferential statistics and results in loss of generality and needs to be dealt with<sup>25</sup>. The values of scale (quadrat size), where the graph of FD versus scale crosses  $FD = 2$ , with respect to the length of the original quadrat size are given in Table 1 (1 unit ~ 30.62 m). Figure 2 shows that the FD cannot be found for the values of scale less than that of the original quadrat size and greater than half of the highest scale used for plotting the semivariogram graphs, leading to lower and upper bounds to the change in quadrat size.

Plots for which ideal sampling scheme can be suggested are listed in Table 2, where  $a$  is the size of side of the quadrat,  $b$  is the minimum distance between two quadrats, and  $c$  is the maximum distance between any two neighbouring quadrats. The value of  $c$ , which is where the graph further increases after the right limb of the  $U$  shape and then starts decreasing sharply, has not been obtained for all the following plots due to constraints on scale.

Figure 3 shows that the plots obtained for species diversity and species richness for a particular forest type are similar, i.e. they have peaks at the same value of

scale. This implies that these two properties for a particular forest type essentially have similar dependence on the scale (distance) transect. Thus FD versus scale plots are also similar for species diversity and species richness for a particular forest type. The semivariograms (Figure 2) of the plots of basal area for all three forests are similar especially basal area plots for SEG and MD forests, inferring that the dependence of basal area on the scale transect does not vary much with the change in forest type. We are not referring to the value of the property ‘basal area’, but to the spatial dependence of the property. The FD versus scale plots show similarities as well; one notable evidence is that all the three basal area plots cross  $FD = 2$ , at two units of scale, which suggest that these values are homogenous quadrats. Semivariograms (Figure 2) of density and height for a particular forest type suggest that both properties vary similarly with the scale transects. This similarity is less when compared to the above-stated cases, especially at higher values of scale.

The suggested quadrat sizes for height and basal area are different implying that even if the particular forest type is defined by these characteristics, they do not vary in the same way with change scale. In a mathematical approach, one can use average values of quadrat sizes suggested for basal area and height.

The fact that FD can be used as a measure of the extent to which a variable is dependent on another, is key for studying variation in properties of vegetation samples along with the fact that FD is not a constant function of scale. One of the primary assumptions of methods of classical inferential statistics is that the samples should be randomly placed or have a homogenous distribution. Selection of quadrat size so that the distribution is homogenous within the quadrat is important to eliminate the pseudo random samples, in order to improve the generality and accuracy of results. FD of 2, which occurs at the suggested value of scale or quadrat size, has a homogenous distribution. By observing the variation of FD along the scale for different properties one can deduce correlations between the spatial dependence of these properties.

1. Evans, T. C. and O’Regan, W. G., Sampling problems in the measurement of range vegetation. In *Range Research Methods*, US Department of Agriculture and Miscellaneous Publications no. 940, 1963, pp. 54–60.
2. Wight, J. R., The sampling unit and its effect on saltbush yield estimates. *J. Range Manage.*, 1967, **20**(5), 323–325.
3. Wilkinson, G. B. and Daly, G., Comparative assessment of some national forest survey types. *N. Z. J. For. Sci.*, 1976, **6**, 363–375.

4. Gillison, A. N. and Brewer, K. R. W., The use of gradient directed transects or gradsects in natural resource surveys. *J. Environ. Manage.*, 1985, **20**, 103–127.
5. Sorrells, L. and Glenn, S., Review of sampling techniques used in studies of grassland plant communities. *Proc. Okla Acad.*, 1991, **71**, 43–45.
6. Jorgensen, E. E., Demarais, S. and Monasmit, T., A variation of line intercept sampling: comparing long transects to short transects. *Texas J. Sci.*, 2000, **52**, 48–52.
7. Allen, R. B. and McLennan, M. J., Indigenous forest survey manual: two inventory methods. Forest Research Institute, New Zealand, FRI Bulletin, 1983, pp. 48–73.
8. Canham, C. D. E., Permanent plotter. Newsletter of the working group on permanent sample plots for the study of vegetation. Institute of Ecosystem Studies, New York Botanic Garden, 1987, p. 6.
9. Sukumar, R., Dattaraja, H., Suresh, H., Radhakrishnan, J., Vasudeva, R., Nirmala, S. and Joshi, N. V., Long-term monitoring of vegetation in a tropical deciduous forest in Mudumalai, southern India. *Curr. Sci.*, 1992, **62**, 608–616.
10. Condit, R., Ashton, P. S., Manokaran, N., LaFrankie, J. V., Hubbell, S. P. and Foster, R. B., Dynamics of the forest communities at Pasoh and Barro Colorado: comparing two 50 ha plots. *Philos. Trans. R. Soc. London, Ser. B*, 1999, **354**, 1739–1748.
11. Laidlaw, M. J., Olsen, M., Kitching, R. I. and Greenway, M., Tree floristic and structural characteristics of one hectare of subtropical rainforests in Lamington National Park, Queensland. *Proc. R. Soc. Queensl.*, 2000, **109**, 91–105.
12. DeWalt, S. J., Bourdy, G. and Chávez-Michel, L. R. and Quenevo C., Ethnobotany of the Tacana: quantitative inventories of two permanent plots of northwestern Bolivia. *Econ. Bot.*, 1999, **53**, 237–260.
13. Valencia, R. *et al.*, Tree species distributions and local habitat variation in the Amazon: a large forest plot in eastern Ecuador. *J. Ecol.*, 2004, **92**, 214–229.
14. Harms, K. E., Condit, R., Hubbell, S. P. and Foster, R. B., Habitat associations of trees and shrubs in a 50-ha neotropical forest plot. *J. Ecol.*, 2001, **89**, 947–959.
15. Parthasarathy, N., Arthur, S. M. and Udayakumar, M., Tropical dry evergreen forests of peninsular India: ecology and conservation significance. *Trop. Conserv. Sci.*, 2008, **1**(2), 89–110.
16. Okland, R. H., Vegetation ecology: theory, methods and applications with reference to Fennoscandia. *Sommerfeltia Suppl.*, 1990, **34**, 1–223.
17. Stewart, G. H., Wardle, J. A. and Burrows, L. E., Forest understorey changes after reduction in deer numbers, North Fiordland, New Zealand. *N. Z. J. Ecol.*, 1987, 1–7.
18. Pascal, J., Ramesh, B. and Franceschi, D., Wet evergreen forest types of the southern Western Ghats, India. *Trop. Ecol.*, 2004, **45**, 281–292.
19. Prasad, P., Sringswara, A., Reddy, C., Kumari, P., Varalakshmi, R., Raza, S. and Dutt, C., Vegetation structure and ecological characteristics of forest of North Andaman Islands. *Biol. Lett.*, 2009, **46**(2), 105–121.
20. Manokaran, N., La Frankie, J. V., Kochummen, K. M., Quah, E. S., Klahn, J. E., Ashton, P. S. and Hubbell, S. P., In *Methodology for the Fifty Hectare Research Plot at Pasoh Forest Reserve* (ed. Chan, H. T.), Forest Research Institute, Malaysia, Research Pamphlet, 1990, vol. 104, pp. 1–69.
21. Parthasarathy, N. and Karthikeyan, R., Biodiversity and population density of woody species in a tropical evergreen forest in Courtallum reserve forest, Western Ghats, India. *Trop. Ecol.*, 1997, **38**(2), 297–306.
22. Lee, H. S., Davies, S. J., LaFrankie, J. V., Tan, S., Yamakura, T., Itoh, A. and Ashton, P. S., Floristic and structural diversity of 52 hectares of mixed dipterocarp forest in Lambir Hills National Park, Sarawak, Malaysia. *J. Trop. For. Sci.*, 2002, **14**, 379–400.
23. Neldner, V. J. and Butler, D. W., Is 500 m<sup>2</sup> an effective plot size to sample floristic diversity for Queensland's vegetation? *Cunninghamia*, 2008, **10**(4), 513–519.
24. Hurlbert, S. H., Pseudoreplication and the design of ecological field experiments. *Ecol. Monogr.*, 1984, **54**(2), 187–211.
25. Palmer, M. W., Fractal geometry: a tool for describing spatial patterns of plant communities. *Vegetation*, 1988, **75**, 91.
26. Bishop, S., *Geographic Information Science and Mountain Geomorphology*, Springer, 2004, pp. 75–97.
27. Cliff, A. D. and Ord, J. K., *Spatial Processes: Models and Applications*, Taylor & Francis, 1981, pp. 23–42.
28. Kadanoff, L. P., Fractals: where's the physics. *Phys. Today*, 1986, **39**, 6.
29. Bak, P., Tang, C. and Wiesenfeld, K., Self-organized criticality. *Phys. Rev. A*, 1988, **38**(1), 364–374.
30. Mandelbrot, B., *The Fractal Geometry of Nature*, W. H. Freeman and Co, New York, 1982.
31. Leduc, A., Prairie, Y. T. and Bergeron, Y., Fractal dimension estimates of a fragmented landscape – sources of variability. *Landsc. Ecol.*, 1994, **9**, 279–286.
32. Horgan, J., Fractal shorthand. *J. Appl. Ecol.*, 1988, **30**, 523–535.
33. Feder, J., *Fractals (Physics of Solids and Liquids)*, Springer, 2003, pp. 655–673.
34. Caroline, M. H., Terry, P. and Richard, T., Fractal dimension of landscape silhouette outlines as a predictor of landscape preference. *J. Environ. Psychol.*, 2004, **24**, 247–255.
35. Stamps, A. E., Fractals, skylines, nature and beauty. *Landsc. Urban Plann.*, 2002, **60**, 163–184.
36. Morse, D. R., Lawton, J. H., Dodson, M. M. and Williamson, M. H., Fractal dimension of vegetation and distribution of arthropod body lengths. *Nature*, 1985, **314**, 731–733.
37. Lorimer, N. D., Haight, R. G. and Leary, R. A., The fractal forest: fractal geometry and applications in forest science. General Technical Report NC-1770, Department of Agriculture, USA, 1994.
38. Falconer, K. J., *Fractal Geometry, Mathematical Foundations and Applications*, John Wiley, Chichester, UK, 1990.
39. Turcotte, D. L., *Fractals and Chaos in Geology and Geophysics*, Cambridge University Press, 1997, p. 2.
40. Kunin, W. E., Extrapolating species abundance across spatial scales. *Science*, 2000, **281**(5382), 1513–1515.
41. Keitt, T. H., Spectral representation of neutral landscapes. *Landsc. Ecol.*, 2000, **15**, 479–493.
42. Halley, J. M. and Iwasa, Y., Neutrality without incoherence: a response to Clark. *Trends Ecol. Evol.*, 2012, **27**(7), 363–363.
43. Garcia, L. S., *The Fractal Explorer*, Dynamic Press, Santa Cruz, USA, 1991.
44. Vicsek, T., *Fractal Growth Phenomena*, World Scientific, Singapore, 1992, vol. 31, pp. 139–146.
45. Falconer, K. J., *The Geometry of Fractal Sets*, Cambridge University Press, 1986.
46. Falconer, K. J., *Fractal Geometry*, Wiley, New York, 2003, p. 308.
47. Burrough, P. A., Multiscale sources of spatial variation in soil. I. The application of fractal concepts to nested levels of soil variation. *J. Soil Sci.*, 1983, **34**, 577–597.
48. Burrough, P. A., Multiscale sources of spatial variation in soil. II. A non-Brownian fractal model and its application in soil survey. *J. Soil Sci.*, 1983, **34**, 599–620.

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